

Chapter 2

General Data Compression Theorems

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Motivations

II: 2-1

- We already know that the entropy rate

$$\lim_{n \rightarrow \infty} \frac{1}{n} H(X^n)$$

is the minimum data compression rate for arbitrarily small data compression error for block coding of the stationary ergodic source.

- We also mentioned that for a more complicated situation where the source becomes non-stationary, the quantity $\lim_{n \rightarrow \infty} (1/n)H(X^n)$ may not exist, and can no longer be used to characterize the source compression.
- This results in the need to establish a new entropy measure which appropriately characterizes the operational limits of arbitrary stochastic systems, which was done in the previous chapter.

Fixed-length codes for arbitrary sources

II: 2-2

- Here, we have made an implicit assumption in the following derivation, which is the source alphabet \mathcal{X} is *finite*.

Definition 2.1 (cf. Definition 3.2 and its associated footnote in [2])

An (n, M) block code for data compression is a set

$$\mathcal{C}_n := \{\mathbf{c}_1, \mathbf{c}_2, \dots, \mathbf{c}_M\}$$

consisting of M sourcewords of block length n (and a binary-indexing codeword for each sourceword \mathbf{c}_i).

Fixed-length codes for arbitrary sources

II: 2-3

Definition 2.2 Fix $\varepsilon \in [0, 1]$. R is an ε -achievable data compression rate for a source \mathbf{X} if there exists a sequence of block data compression codes $\{\mathcal{C}_n = (n, M_n)\}_{n=1}^{\infty}$ with

$$\limsup_{n \rightarrow \infty} \frac{1}{n} \log M_n \leq R,$$

and

$$\limsup_{n \rightarrow \infty} P_e(\mathcal{C}_n) \leq \varepsilon,$$

where $P_e(\mathcal{C}_n) := \Pr(X^n \notin \mathcal{C}_n)$ is the probability of decoding error.

The infimum of all ε -achievable data compression rate for \mathbf{X} is denoted by $T_\varepsilon(\mathbf{X})$.

- Note that in conventional source coding theorem, one wants to find the minimum rate with arbitrary small error. This rate is exactly $\lim_{\varepsilon \downarrow 0} T_\varepsilon(\mathbf{X})$.
- As expected, for DMS, $\lim_{\varepsilon \downarrow 0} T_\varepsilon(\mathbf{X}) = H(X)$. Actually, for DMS, $T_\varepsilon(\mathbf{X}) = H(X)$ for any $\varepsilon \in [0, 1)$.

Fixed-length codes for arbitrary sources

II: 2-4

Lemma 2.3 Fix a positive integer n . There exists an (n, M_n) source block code \mathcal{C}_n for P_{X^n} such that its error probability satisfies

$$P_e(\mathcal{C}_n) \leq \Pr \left[\frac{1}{n} h_{X^n}(X^n) > \frac{1}{n} \log M_n \right].$$

Proof: Observe that

$$\begin{aligned} 1 &\geq \sum_{\{x^n \in \mathcal{X}^n : (1/n)h_{X^n}(x^n) \leq (1/n) \log M_n\}} P_{X^n}(x^n) \\ &\geq \sum_{\{x^n \in \mathcal{X}^n : (1/n)h_{X^n}(x^n) \leq (1/n) \log M_n\}} \frac{1}{M_n} \\ &\geq \left| \left\{ x^n \in \mathcal{X}^n : \frac{1}{n} h_{X^n}(x^n) \leq \frac{1}{n} \log M_n \right\} \right| \frac{1}{M_n}. \end{aligned}$$

Therefore, $|\{x^n \in \mathcal{X}^n : (1/n)h_{X^n}(x^n) \leq (1/n) \log M_n\}| \leq M_n$. We can then choose a code

$$\mathcal{C}_n \supset \left\{ x^n \in \mathcal{X}^n : \frac{1}{n} h_{X^n}(x^n) \leq \frac{1}{n} \log M_n \right\}$$

with $|\mathcal{C}_n| = M_n$ and

$$P_e(\mathcal{C}_n) = 1 - P_{X^n}\{\mathcal{C}_n\} \leq \Pr \left[\frac{1}{n} h_{X^n}(X^n) > \frac{1}{n} \log M_n \right].$$

□

Fixed-length codes for arbitrary sources

II: 2-5

Lemma 2.4 Every (n, M_n) source block code \mathcal{C}_n for P_{X^n} satisfies

$$P_e(\mathcal{C}_n) \geq \Pr \left[\frac{1}{n} h_{X^n}(X^n) > \frac{1}{n} \log M_n + \gamma \right] - \exp\{-n\gamma\},$$

for every $\gamma > 0$.

Proof: It suffices to prove that

$$1 - P_e(\mathcal{C}_n) = \Pr \{X^n \in \mathcal{C}_n\} \leq \Pr \left[\frac{1}{n} h_{X^n}(X^n) \leq \frac{1}{n} \log M_n + \gamma \right] + \exp\{-n\gamma\}.$$

Clearly,

$$\begin{aligned} \Pr \{X^n \in \mathcal{C}_n\} &= \Pr \left\{ X^n \in \mathcal{C}_n \text{ and } \frac{1}{n} h_{X^n}(X^n) \leq \frac{1}{n} \log M_n + \gamma \right\} \\ &\quad + \Pr \left\{ X^n \in \mathcal{C}_n \text{ and } \frac{1}{n} h_{X^n}(X^n) > \frac{1}{n} \log M_n + \gamma \right\} \end{aligned}$$

Fixed-length codes for arbitrary sources

II: 2-6

$$\begin{aligned}
 &\leq \Pr \left\{ \frac{1}{n} h_{X^n}(X^n) \leq \frac{1}{n} \log M_n + \gamma \right\} \\
 &\quad + \Pr \left\{ X^n \in \mathcal{C}_n \text{ and } \frac{1}{n} h_{X^n}(X^n) > \frac{1}{n} \log M_n + \gamma \right\} \\
 &= \Pr \left\{ \frac{1}{n} h_{X^n}(X^n) \leq \frac{1}{n} \log M_n + \gamma \right\} \\
 &\quad + \sum_{x^n \in \mathcal{C}_n} P_{X^n}(x^n) \cdot \mathbf{1} \left\{ \frac{1}{n} h_{X^n}(x^n) > \frac{1}{n} \log M_n + \gamma \right\} \\
 &= \Pr \left\{ \frac{1}{n} h_{X^n}(X^n) \leq \frac{1}{n} \log M_n + \gamma \right\} \\
 &\quad + \sum_{x^n \in \mathcal{C}_n} P_{X^n}(x^n) \cdot \mathbf{1} \left\{ P_{X^n}(x^n) < \frac{1}{M_n} \exp\{-n\gamma\} \right\} \\
 &< \Pr \left\{ \frac{1}{n} h_{X^n}(X^n) \leq \frac{1}{n} \log M_n + \gamma \right\} + |\mathcal{C}_n| \frac{1}{M_n} \exp\{-n\gamma\} \\
 &= \Pr \left\{ \frac{1}{n} h_{X^n}(X^n) \leq \frac{1}{n} \log M_n + \gamma \right\} + \exp\{-n\gamma\}.
 \end{aligned}$$

□

Fixed-length codes for arbitrary sources

II: 2-7

We now apply Lemmas 2.3 and 2.4 to prove a *general* source coding theorem for block codes.

Theorem 2.5 (general source coding theorem) For any source \mathbf{X} ,

$$T_\varepsilon(\mathbf{X}) = \begin{cases} \lim_{\delta \uparrow (1-\varepsilon)} \bar{H}_\delta(\mathbf{X}), & \text{for } \varepsilon \in [0, 1); \\ 0, & \text{for } \varepsilon = 1. \end{cases}$$

Proof: The case of $\varepsilon = 1$ follows directly from its definition; hence, the proof only focus on the case of $\varepsilon \in [0, 1)$.

1. . *Forward part (achievability):* $T_\varepsilon(\mathbf{X}) \leq \lim_{\delta \uparrow (1-\varepsilon)} \bar{H}_\delta(\mathbf{X})$

We need to prove the existence of a sequence of block codes $\{\mathcal{C}_n = (n, M_n)\}_{n \geq 1}$ such that for every $\gamma > 0$,

$$\limsup_{n \rightarrow \infty} \frac{1}{n} \log M_n \leq \lim_{\delta \uparrow (1-\varepsilon)} \bar{H}_\delta(\mathbf{X}) + \gamma \quad \text{and} \quad \limsup_{n \rightarrow \infty} P_e(\mathcal{C}_n) \leq \varepsilon.$$

Lemma 2.3 ensures the existence (for any $\gamma > 0$) of a source block code $\mathcal{C}_n = (n, M_n = \lceil \exp\{n(\lim_{\delta \uparrow (1-\varepsilon)} \bar{H}_\delta(\mathbf{X}) + \gamma)\} \rceil)$ with error probability

$$\begin{aligned} P_e(\mathcal{C}_n) &\leq \Pr \left\{ \frac{1}{n} h_{X^n}(X^n) > \frac{1}{n} \log M_n \right\} \\ &\leq \Pr \left\{ \frac{1}{n} h_{X^n}(X^n) > \lim_{\delta \uparrow (1-\varepsilon)} \bar{H}_\delta(\mathbf{X}) + \gamma \right\}. \end{aligned}$$

Therefore,

$$\begin{aligned}
 \limsup_{n \rightarrow \infty} P_e(\mathcal{C}_n) &\leq \limsup_{n \rightarrow \infty} \Pr \left\{ \frac{1}{n} h_{X^n}(X^n) > \lim_{\delta \uparrow (1-\varepsilon)} \bar{H}_\delta(\mathbf{X}) + \gamma \right\} \\
 &= 1 - \liminf_{n \rightarrow \infty} \Pr \left\{ \frac{1}{n} h_{X^n}(X^n) \leq \lim_{\delta \uparrow (1-\varepsilon)} \bar{H}_\delta(\mathbf{X}) + \gamma \right\} \\
 &\leq 1 - (1 - \varepsilon) = \varepsilon,
 \end{aligned}$$

where the last inequality follows from

$$\lim_{\delta \uparrow (1-\varepsilon)} \bar{H}_\delta(\mathbf{X}) = \sup \left\{ \theta : \liminf_{n \rightarrow \infty} \Pr \left[\frac{1}{n} h_{X^n}(X^n) \leq \theta \right] < 1 - \varepsilon \right\}. \quad (2.1.1)$$

2. *Converse part:* $T_\varepsilon(\mathbf{X}) \geq \lim_{\delta \uparrow (1-\varepsilon)} \bar{H}_\delta(\mathbf{X})$

Assume without loss of generality that $\lim_{\delta \uparrow (1-\varepsilon)} \bar{H}_\delta(\mathbf{X}) > 0$. We will prove the converse by contradiction. Suppose that $T_\varepsilon(\mathbf{X}) < \lim_{\delta \uparrow (1-\varepsilon)} \bar{H}_\delta(\mathbf{X})$. Then $(\exists \gamma > 0)$ $T_\varepsilon(\mathbf{X}) < \lim_{\delta \uparrow (1-\varepsilon)} \bar{H}_\delta(\mathbf{X}) - 4\gamma$. By definition of $T_\varepsilon(\mathbf{X})$, there exists a sequence of codes $\mathcal{C}_n = (n, M_n)$ such that

$$\limsup_{n \rightarrow \infty} \frac{1}{n} \log M_n \leq \left(\lim_{\delta \uparrow (1-\varepsilon)} \bar{H}_\delta(\mathbf{X}) - 4\gamma \right) + \gamma < \lim_{\delta \uparrow (1-\varepsilon)} \bar{H}_\delta(\mathbf{X}) - 2\gamma \quad (2.1.2)$$

and

$$\limsup_{n \rightarrow \infty} P_e(\mathcal{C}_n) \leq \varepsilon. \quad (2.1.3)$$

Fixed-length codes for arbitrary sources

II: 2-9

(2.1.2) implies that

$$\frac{1}{n} \log M_n \leq \lim_{\delta \uparrow (1-\varepsilon)} \bar{H}_\delta(\mathbf{X}) - 2\gamma$$

for all sufficiently large n . Hence, for those n satisfying the above inequality and also by Lemma 2.4,

$$\begin{aligned} P_e(\mathcal{C}_n) &\geq \Pr \left[\frac{1}{n} h_{X^n}(X^n) > \frac{1}{n} \log M_n + \gamma \right] - e^{-n\gamma} \\ &\geq \Pr \left[\frac{1}{n} h_{X^n}(X^n) > \left(\lim_{\delta \uparrow (1-\varepsilon)} \bar{H}_\delta(\mathbf{X}) - 2\gamma \right) + \gamma \right] - e^{-n\gamma}. \end{aligned}$$

Therefore,

$$\begin{aligned} \limsup_{n \rightarrow \infty} P_e(\mathcal{C}_n) &\geq 1 - \liminf_{n \rightarrow \infty} \Pr \left[\frac{1}{n} h_{X^n}(X^n) \leq \lim_{\delta \uparrow (1-\varepsilon)} \bar{H}_\delta(\mathbf{X}) - \gamma \right] \\ &> 1 - (1 - \varepsilon) = \varepsilon, \end{aligned}$$

where the last inequality follows from (2.1.1). Thus, a contradiction to (2.1.3) is obtained. \square

Fixed-length codes for arbitrary sources

II: 2-10

A few remarks are made based on the previous theorem.

- Note that as $\varepsilon \rightarrow 0$,

$$T_0(\mathbf{X}) = \lim_{\delta \uparrow (1-\varepsilon)} \bar{H}_\delta(\mathbf{X}) = \bar{H}(\mathbf{X}).$$

Hence, the minimum (asymptotic) lossless fixed-length source coding rate of any finite-alphabet source is $\bar{H}(\mathbf{X})$.

- Consider the special case where

$$-\frac{1}{n} \log P_{X^n}(X^n) \text{ converges in probability to a constant } H \text{ (entropy rate),}$$

which holds for all *information stable* sources. In this case, both the inf- and sup-spectrums of \mathbf{X} degenerate to a unit step function:

$$u(\theta) = \begin{cases} 1, & \text{if } \theta > H; \\ 0, & \text{if } \theta < H. \end{cases}$$

Thus, $\bar{H}_\varepsilon(\mathbf{X}) = H$ for all $\varepsilon \in [0, 1)$. Hence, general source coding theorem reduces to the conventional source coding theorem.

Fixed-length codes for arbitrary sources

II: 2-11

- A source

$$\mathbf{X} = \left\{ X^n = \left(X_1^{(n)}, \dots, X_n^{(n)} \right) \right\}_{n=1}^{\infty}$$

is said to be *information stable* if

$$H(X^n) = E[-\log P_{X^n}(x^n)] > 0 \text{ for all } n,$$

and

$$\lim_{n \rightarrow \infty} \Pr \left(\left| \frac{-\log P_{X^n}(x^n)}{H(X^n)} - 1 \right| > \varepsilon \right) = 0,$$

for every $\varepsilon > 0$.

- By the definition, any stationary-ergodic source with finite n -fold entropy is information stable; hence, it can be viewed as a generalized source model for stationary-ergodic sources.

Fixed-length codes for arbitrary sources

II: 2-12

- If

$-\frac{1}{n} \log P_{X^n}(X^n)$ converges in probability to a random variable Z

whose cdf is $F_Z(\cdot)$, then the minimum achievable data compression rate subject to decoding error being no greater than ε is

$$T_\varepsilon(\mathbf{X}) = \lim_{\delta \uparrow (1-\varepsilon)} \bar{H}_\delta(\mathbf{X}) = \sup \{R : F_Z(R) < 1 - \varepsilon\}.$$

Example 2.6 Consider a binary source \mathbf{X} with each X^n is Bernoulli(Θ) distributed, where Θ is a random variable defined over $(0, 1)$. By ergodic decomposition theorem (which states that any stationary source can be viewed as a mixture of stationary-ergodic sources) that

$-\frac{1}{n} \log P_{X^n}(X^n)$ converges in probability to $h_b(\Theta)$,

where $h_b(x) := -x \log_2(x) - (1-x) \log_2(1-x)$. Consequently,

$$T_\varepsilon(\mathbf{X}) = \sup \{R : \Pr\{h_b(\Theta) \leq R\} < 1 - \varepsilon\}.$$

Fixed-length codes for arbitrary sources

II: 2-13

- From the above example, or from Theorem 2.5, it shows that the *strong converse theorem* (which states that codes with rate below entropy rate will ultimately have decompression error approaching one) does not hold in general. However, one can always claim the *weak converse statement* for arbitrary sources.

Theorem 2.7 (weak converse theorem) For any block code sequence of ultimate rate $R < \bar{H}(\mathbf{X})$, the probability of block decoding failure P_e cannot be made arbitrarily small. In other words, there exists $\varepsilon > 0$ such that P_e is lower bounded by ε infinitely often in block length n .

$$\begin{array}{c}
 P_e \xrightarrow{n \text{ (i.o.)}} 1 \quad \left| \quad P_e \text{ is lower} \right. \quad P_e \xrightarrow{n \rightarrow \infty} 0 \\
 \left. \text{bounded (i.o. in } n \text{)} \right| \\
 \hline
 \bar{H}_0(\mathbf{X}) \qquad \qquad \qquad \bar{H}(\mathbf{X}) \qquad \qquad \qquad R
 \end{array}$$

Behavior of the probability of block decoding error as block length n goes to infinity for an *arbitrary* source \mathbf{X} .

Generalized AEP theorem

II: 2-14

Theorem 2.8 (generalized asymptotic equipartition property for arbitrary sources) Fix $\varepsilon \in [0, 1)$. Given an arbitrary source \mathbf{X} , define

$$\mathcal{T}_n[R] := \left\{ x^n \in \mathcal{X}^n : -\frac{1}{n} \log P_{X^n}(x^n) \leq R \right\}.$$

Then for any $\delta > 0$, the following statements hold.

1.

$$\liminf_{n \rightarrow \infty} \Pr \left\{ \mathcal{T}_n[\bar{H}_\varepsilon(\mathbf{X}) - \delta] \right\} \leq \varepsilon \quad (2.1.4)$$

2.

$$\liminf_{n \rightarrow \infty} \Pr \left\{ \mathcal{T}_n[\bar{H}_\varepsilon(\mathbf{X}) + \delta] \right\} > \varepsilon \quad (2.1.5)$$

3. The number of elements in

$$\mathcal{F}_n(\delta; \varepsilon) := \mathcal{T}_n[\bar{H}_\varepsilon(\mathbf{X}) + \delta] \setminus \mathcal{T}_n[\bar{H}_\varepsilon(\mathbf{X}) - \delta],$$

denoted by $|\mathcal{F}_n(\delta; \varepsilon)|$, satisfies

$$|\mathcal{F}_n(\delta; \varepsilon)| \leq \exp \left\{ n(\bar{H}_\varepsilon(\mathbf{X}) + \delta) \right\}, \quad (2.1.6)$$

where the operation $\mathcal{A} \setminus \mathcal{B}$ between two sets \mathcal{A} and \mathcal{B} is defined by $\mathcal{A} \setminus \mathcal{B} := \mathcal{A} \cap \mathcal{B}^c$ with \mathcal{B}^c denoting the complement set of \mathcal{B} .

Generalized AEP theorem

II: 2-15

4. There exists $\rho = \rho(\delta) > 0$ and a subsequence $\{n_j\}_{j=1}^{\infty}$ such that

$$|\mathcal{F}_n(\delta; \varepsilon)| > \rho \cdot \exp \{n_j(\bar{H}_\varepsilon(\mathbf{X}) - \delta)\}. \quad (2.1.7)$$

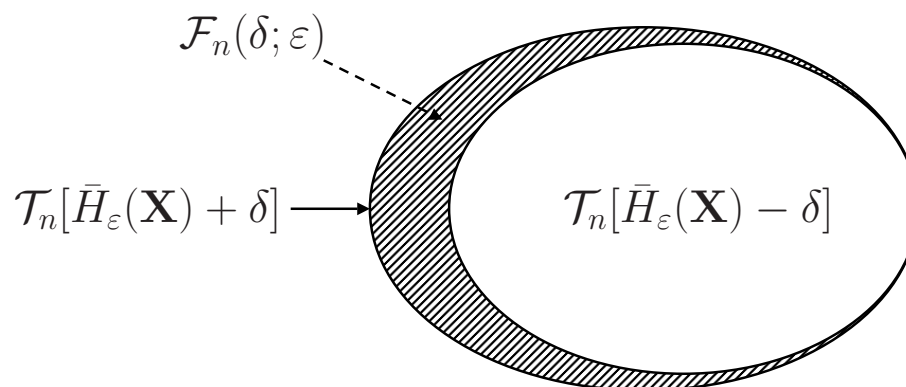


Illustration of generalized AEP Theorem. $\mathcal{F}_n(\delta; \varepsilon) := \mathcal{T}_n[\bar{H}_\varepsilon(\mathbf{X}) + \delta] \setminus \mathcal{T}_n[\bar{H}_\varepsilon(\mathbf{X}) - \delta]$ is the dashed region.

Generalized AEP theorem

II: 2-16

- The set

$$\begin{aligned} \mathcal{F}_n(\delta; \varepsilon) &:= \mathcal{T}_n[\bar{H}_\varepsilon(\mathbf{X}) + \delta] \setminus \mathcal{T}_n[\bar{H}_\varepsilon(\mathbf{X}) - \delta] \\ &= \left\{ x^n \in \mathcal{X}^n : \left| -\frac{1}{n} \log P_{X^n}(x^n) - \bar{H}_\varepsilon(\mathbf{X}) \right| < \delta \right\} \\ &\quad \cup \left\{ x^n \in \mathcal{X}^n : -\frac{1}{n} \log P_{X^n}(x^n) = \bar{H}_\varepsilon(\mathbf{X}) + \delta \right\} \end{aligned}$$

is nothing but the weakly δ -typical set.

- $q_n := \Pr\{\mathcal{F}_n(\delta; \varepsilon)\} > 0$ infinitely often in n .

-

$$|\mathcal{F}_n(\delta; \varepsilon)| \approx e^{n\bar{H}_\varepsilon(\mathbf{X})},$$

and the probability of each sequence in $\mathcal{F}_n(\delta; \varepsilon)$ can be estimated by $q_n \cdot \exp\{-n\bar{H}_\varepsilon(\mathbf{X})\}$.

- In particular, if \mathbf{X} is a stationary-ergodic source, then $\bar{H}_\varepsilon(\mathbf{X})$ is independent of $\varepsilon \in [0, 1)$ and, $\bar{H}_\varepsilon(\mathbf{X}) = \underline{H}_\varepsilon(\mathbf{X}) = H$ for all $\varepsilon \in [0, 1)$, where H is the source entropy rate

$$H = \lim_{n \rightarrow \infty} \frac{1}{n} E[-\log P_{X^n}(X^n)].$$

In this case, the generalized AEP reduces to the conventional AEP.